

Species Profiles Support Recommendations for Quality Filtering of Opportunistic Citizen Science Data

Camille Van Eupen^{*,a,b}, Dirk Maes^{c,d}, Marc Herremans^e, Kristijn R.R. Swinnen^e, Ben Somers^{**b} and Stijn Luca^{**a}

* corresponding author

** joint last authors

^a Ghent University, Department of Data Analysis and Mathematical Modelling, Coupure Links 653, B-9000 Ghent, Belgium; camille.vaneupen@kuleuven.be (ORCID 0000-0002-0924-8892); stijn.luca@ugent.be (ORCID 0000-0002-6781-7870)

^b KU Leuven, Department of Earth and Environmental Sciences, Division Forest Nature and Landscape, Celestijnenlaan 200E, B-3001 Heverlee, Belgium; ben.somers@kuleuven.be (ORCID 0000-0002-7875-107X)

^c Research Institute for Nature and Forest (INBO), Herman Teirlinckgebouw, Havenlaan 88 box 73, B-1000 Brussels, Belgium; dirk.maes@inbo.be (ORCID 0000-0002-7947-3788)

^d Radboud Institute for Biological and Environmental Sciences (RIBES), Radboud University, PO Box 9010, NL-6500 GL Nijmegen, The Netherlands; dirk.maes@ru.nl

^e Natuurpunt Studie, Coxiestraat 11, 2800 Mechelen, Belgium; ; marc.herremans@natuurpunt.be (ORCID: 0000-0002-9719-8732); kristijn.swinnen@natuurpunt.be (ORCID 0000-0002-1910-9247)

1 **ABSTRACT**

2 Opportunistic citizen science data are commonly filtered in an attempt to improve their applicability for
3 relating species occurrences with environmental variables. Recommendations on when and how to filter,
4 however, have remained relatively general and associations between species traits and filtering
5 recommendations are sparse. We collected six traits (body size, detectability, classification error rate,
6 familiarity, reporting probability and range size) of 52 birds, 25 butterflies and 14 dragonflies. Both
7 absolute (values not rescaled) and relative traits (values rescaled per taxonomic group) were linked to
8 filter effects, i.e. the impact on three different measures of species distribution model performance
9 caused by applying three different quality filters, for different degrees of sample size reduction. First, we
10 applied multiple regressions that predicted the filter effects by either absolute (including taxonomic
11 group) or relative traits. Second, a principal component and clustering analysis were performed to define
12 five species profiles based on species traits that were retained after a multiple regression model selection.
13 The analysis of the profiles indicated the relative importance of species traits and revealed new insights
14 into the association of species traits with changes in model performance after data quality filtering. Both
15 taxonomic group (more than absolute traits) and relative species traits (mainly classification error rate,
16 range size and familiarity) defined the impact of data quality filtering on model performance and we
17 discourage the selection of a quality filtering strategy based on one single species trait. Results further
18 confirmed the importance of considering the goal of the study (i.e. increasing model discrimination
19 capacity, sensitivity or specificity) as well as the change in sample size caused by stringent filtering. The
20 general species knowledge among citizen scientists (importance of observer experience), together with
21 the mechanism of record verification in an opportunistic data platform (importance of verifiable
22 metadata) have the largest potential for enhancing the quality of opportunistic records.

23 **KEYWORDS**

- 24 data quality filtering, filtering recommendations, opportunistic data, presence-only data, species
- 25 distribution models, species traits

26 **1. Introduction**

27 Biodiversity conservation needs adequate monitoring of species (Lindenmayer et al., 2020), especially in
28 times of rapid changes to the environment that threaten species and reduce abundances at alarming rates
29 (Newbold et al., 2015; Urban et al., 2016). Structured surveys, where species are recorded in a
30 standardized manner, are commonly put forward as the most desirable strategy for biodiversity
31 monitoring because of their high information content (Dobson et al., 2020). It is challenging, however, to
32 organize structured surveys for a large variety of species over broad spatial and temporal scales, leading
33 to spatial and temporal data gaps (Urban et al., 2016). When information on the potential distributions of
34 species is needed, e.g. for assigning protected areas (Thomaes et al., 2008) or areas of high potential
35 nature value (Maes et al., 2005), spatial gaps can be filled by using species distribution models (SDMs).
36 These models link environmental parameters, such as landscape and climate variables, to species
37 occurrence records (Guisan and Zimmermann, 2000).

38 Species occurrence records with high information content (e.g. collected in structured surveys) are
39 preferably used as input for SDMs, but as aforementioned, such data are sparse. Therefore, SDMs are
40 increasingly built with data collected by citizen scientists in a semi-structured manner (termed semi-
41 structured data) or in an opportunistic unstructured manner (termed opportunistic data). Large citizen
42 science initiatives that have online data platforms either focus on one data type (e.g. eBird contains only
43 semi-structured data (Sullivan et al., 2009), iNaturalist contains only opportunistic data
44 (<https://www.inaturalist.org/>)) or they combine both data types (e.g. iRecord
45 (<https://www.brc.ac.uk/irecord/>), waarnemingen.be (<https://www.waarnemingen.be>) and
46 Observation.org (<https://observation.org>)). The majority of opportunistic data consist of species presence
47 records with some basic information such as the date and geographical precision of the observation
48 (termed opportunistic presence-only data). The main advantage of opportunistic data is the availability in
49 large amounts, over large geographical areas and potentially long periods (Kosmala et al., 2016). A major

50 disadvantage, however, is the prevalence of different types of bias and error (Bird et al., 2014; Isaac and
51 Pocock, 2015) caused by a lack of standardised design and accompanying metadata. Opportunistic citizen
52 science data is therefore associated with uncertainty and scepticism towards its use for SDMs (Burgess et
53 al., 2017), and has led to many publications on how and when to use this type of data for biodiversity
54 research (e.g. Henckel et al., 2020; Isaac & Pocock, 2015; Maes et al., 2015; Van Strien et al., 2013). Ideally,
55 the benefits of both unstructured opportunistic data (quantity) and (semi-)structured survey data
56 (information content) are exploited simultaneously. This can be in the form of model-based data
57 integration (for a recent review see Isaac et al., 2020) or when structured data is used as external
58 validation data for SDMs (e.g. Matutini et al., 2021; Van Eupen et al., 2021).

59 In preparation for an SDM study, data are expected to be cleansed (Zurell et al., 2020). At a minimum, this
60 includes the removal of spatial and temporal outliers, duplicates and records with low precision (Serra-
61 Diaz et al., 2017). When dealing with opportunistic data, cleansing usually also implies stringent filtering,
62 where data are filtered based on record attributes that hold information on the observation process or
63 post-entry data validation (Steen et al., 2019). Even though drawing direct ecological inferences from
64 opportunistic observations is not recommended (Dobson et al., 2020), clear recommendations on when
65 and how to filter opportunistic data remain sparse (but see e.g. Kamp et al., 2016; Steen et al., 2019; Van
66 Eupen et al., 2021; Vantieghem et al., 2017). The study of Van Eupen et al. (2021) highlighted the
67 importance of considering both the type of filter and the resulting change in sample size, yet variation
68 among species in their response to data quality filtering remained large. Grouping species according to
69 their taxonomy revealed that filtering benefitted some groups (i.e. plants and dragonflies) more than
70 others (i.e. butterflies and birds). In this paper, we aim to verify whether grouping species according to a
71 priori selected life-history and/or ecological traits could better substantiate recommendations for data
72 quality filtering.

73 Species traits have been linked extensively to SDM performance and those that cause most variation can
74 usually (but not exhaustively) be compiled to the following three: (1) traits that define the species-
75 environment relationship (e.g. range size, niche breadth (Brotons et al., 2007; Stockwell and Peterson,
76 2002) and habitat association (Chefaoui et al., 2011)), (2) traits that impact the detectability of the species
77 in space and time (e.g. conspicuousness (Seoane et al., 2005), migratory behaviour (Carrascal et al., 2006)
78 and lifespan (Hanspach et al., 2010)), and (3) traits that influence the proneness to misidentification (e.g.
79 phylogenetic relatedness (Vantieghem et al., 2017)).

80 Notwithstanding the vast amount of proof on the link between species traits and absolute SDM
81 performance, few studies have successfully linked species traits to the change in SDM performance caused
82 by stringent filtering of species occurrence records (but see e.g. Steen et al., 2019, where models of more
83 restricted species performed better when using data collected with lower effort). This could be due to the
84 higher quality of the unfiltered data in most of these studies (e.g. semi-structured data in Steen et al.
85 (2019)) or due to the conflicting character of the simultaneous impact of data quality filtering, i.e. an
86 increase in data quality and a decrease in sample size (Van Eupen et al., 2021). By assessing this twofold
87 effect on an extensive dataset of opportunistic records, *waarnemingen.be*, we will aid the optimisation of
88 the data cleansing process that is essential for high-quality SDMs (Zurell et al., 2020).

89 **2. Material and methods**

90 **2.1. Species data and impact of quality filtering**

91 We used a dataset from a previous study on data quality filtering (Van Eupen et al., 2021, Van Eupen et
92 al., 2021b), where three dichotomous filters were applied to opportunistic species observations belonging
93 to four well-studied taxonomic groups in Flanders, i.e. birds, butterflies, dragonflies and plants. Plant
94 observations were not used in the present analysis because their traits are not directly comparable to
95 animal species traits. Data were collected from the '*waarnemingen.be*' database, an online citizen science
96 platform that contains both semi-structured and opportunistic records (Swinnen et al., n.d.). The selected
97 dataset included year-round records from January 2014 to September 2019 with a geographical precision
98 of 500 metres or smaller (records can be submitted to the platform as point locations with specified
99 precision or as observations made within a larger area). Potential seasonal changes in habitat occupancy
100 or range size were not considered. The selection excluded records verified as incorrect by species experts
101 or the data platform's auto-validation system (Swinnen et al., 2018) and records from non-native or non-
102 breeding birds in Flanders (Vermeersch et al., 2020). Opportunistic presence-only data were used for
103 model training, also excluding absences (zero-counts), and semi-structured data for model testing (Van
104 Eupen et al., 2021). The three filters were: '**ACTIVITY**', based on an observer's average annual activity rate,
105 where the filter consists in removing records from less active observers; '**DETAIL**', based on the presence
106 of metadata beyond default requirements (i.e. species name, location, date and observer id), where the
107 filter consists in removing records that were submitted without any additional information (e.g. sex,
108 count, behaviour); and '**VALSTAT**', based on the validation status of a record in the data platform, where
109 the filter consists in removing doubtful and unevaluated records (Table 1). These are all records that could
110 not be verified by species experts because key information was missing or because the record was not
111 assessed yet by an expert at the moment the dataset was extracted.

112 *Table 1: Overview and definitions of the used variables in this study.*

| Data quality filters | <i>description:</i> | <i>based on:</i> |
|--|--|---|
| ACTIVITY | removes records from less active observers | an observer's average annual activity rate |
| DETAIL | removes records that were submitted without any additional information | the presence of metadata beyond default requirements |
| VALSTAT | removes doubtful and unevaluated records | the validation status of a record in the data platform |
| Species traits | <i>description:</i> | <i>source:</i> |
| Body size | wing length (birds and butterflies) or head-to-tail length (dragonflies) | Bink (1992); Storchová and Hořák (2018), https://www.vlinderstichting.nl/libellen/ |
| Classification error rate | the number of erroneous photo records (i.e. observations accompanied by a photograph) relative to the total number of photo records. | the <i>waarnemingen.be</i> data portal during the study period |
| Detectability | the probability of detecting a species on the condition that it is present | quantified by applying site occupancy models to complete checklist data, retrieved from the <i>waarnemingen.be</i> data portal |
| Familiarity | reflects how well-known a species is by the average observer | number of search results retrieved from the Google search engine |
| Reporting probability | the likelihood that a species is reported by an average observer, on the condition that it is present and that the taxonomic group it belongs to is surveyed | a species' relative (per taxonomic group) average reporting rate divided by its detectability, retrieved from the <i>waarnemingen.be</i> data portal |
| Range size | the distribution range size | the total number of grid cells (km ²) in which a species has been recorded during the study period, retrieved from the <i>waarnemingen.be</i> data portal |
| Absolute traits | unscaled trait values as retrieved by the different methods described | |
| Relative traits | scaled trait values; using the following transformation per taxonomic group: $y = \frac{x - \min(x)}{\max(x) - \min(x)}$ | |
| Impact on model performance | | |
| Δ AUC | change in the area under the receiver operating characteristic | |
| Δ sensitivity | change in the true positive rate (TPR) after data quality filtering $TPR = \frac{\text{true positives}}{\text{true positives} + \text{false negatives}}$ | |
| Δ specificity | change in the true negative rate (TNR) after data quality filtering $TNR = \frac{\text{true negatives}}{\text{true negatives} + \text{false positives}}$ | |
| Filter effects | | |
| All combinations of data quality filters and impact on model performance | the impact of data quality filtering (Δ AUC, Δ sensitivity and Δ specificity) by the three filters ACTIVITY, DETAIL and VALSTAT | |
| Sample size situations | | |
| real | the actual reduction in the number of presences after data quality filtering | |
| r50 | a relative reduction in the number of presences after data quality filtering of more than 50% | |
| ss100 | a reduction to 100 presences after data quality filtering | |

113 After filtering, records were aggregated to a 1x1 km resolution to reduce spatial bias (Kramer-Schadt et
114 al., 2013). Species were selected based on opportunistic data availability for model training as well as
115 semi-structured data availability for model validation (Van Eupen et al., 2021). The impact of filtering on
116 the performance of Maxent (Phillips et al., 2006) was assessed by evaluating the difference in three
117 commonly used evaluation metrics of model discrimination before and after filtering: the area under the
118 receiver operating characteristic (AUC), sensitivity and specificity (Fielding and Bell, 1997). Their use was
119 justified because models were run for the same geographical extent and model predictions were
120 evaluated on a testing set that contained the same presences and absences per species (Jiménez-
121 Valverde, 2012; Lobo et al., 2008). Model performance could therefore be interpreted in a relative
122 manner, where an increase in AUC after filtering implied that the filtered data produced models that could
123 better distinguish between testing presences and absences, and increases in sensitivity and specificity
124 implied a higher predicted positive and negative fraction respectively. For the analysis in this study, we
125 extracted the change in model performance (i.e. Δ AUC, Δ sensitivity and Δ specificity) (Table 1), after
126 using the three single filters (ACTIVITY, DETAIL and VALSTAT) for 52 birds, 25 butterflies and 14
127 dragonflies. For a summary per species of the data used for model testing and model training (unfiltered
128 and filtered data) and of the impact on model performance, we refer to the supplementary information
129 1 (Table C.1) and supplementary information 2 respectively in the study of Van Eupen et al. (2021).

130 **2.2. Species traits**

131 We used six species traits that can be related to data quality in opportunistic citizen science data based
132 on literature review and expert opinion: body size, detectability, classification error rate, familiarity,
133 reporting probability and range size (Table 1). Abundance was not considered because the largely
134 unstructured *waarnemingen.be* database contains unreliable count data that are mostly without a clear
135 reference of time and space. All trait values can be found in the supplementary information (Table S1).

136 **Body size** equals the wing length for birds (Storchová and Hořák, 2018) and butterflies (Bink, 1992) and
137 head-to-tail length for dragonflies (<https://www.vlinderstichting.nl/libellen/>).

138 The **classification error rate** reflects how likely it is for an average observer to wrongly identify a species.
139 This was quantified by the number of erroneous photo records (i.e. observations accompanied by a
140 photograph) of a species in the *waarnemingen.be* data portal, relative to its total number of photo
141 records. The portal keeps track of changes in the identification of a species, and we considered only the
142 changes at the species-level as erroneous (and for example not the changes from family or genus to
143 species-level). Auto-corrections made by the observer were excluded.

144 **Detectability** is the probability of detecting a species on the condition that it is present (MacKenzie et al.,
145 2017). Species detectability was retrieved from applying site occupancy models to complete semi-
146 structured checklist data extracted from *waarnemingen.be*, following Johnston et al. (2021). Detection
147 histories consisted of five to ten repeated visits to a specific site (a 1 km grid cell) by the same observer in
148 a period of closure (i.e. a period with no supposed changes in occupancy). A period of closure was defined
149 as 20 consecutive days in the peak active season of a species. The peak active season was defined as every
150 10 days with an observation count above the average count of all observations in a year, excluding egg,
151 larva, pupa and caterpillar observations. Covariates used to describe the detection process were: checklist
152 duration (in minutes), starting time of the checklist, search effort (i.e. the number of species recorded at
153 a specific location, supporting on the principle of species accumulation curves (Colwell et al., 2004)), and
154 open habitat (grasslands, wetland, marshes and water) versus closed habitat (forest and woodland),
155 because of an increased detectability (visually and, for birds, also auditory) in open habitat types (Johnston
156 et al., 2014; Morton, 1975). Detection probabilities were predicted for all grids with covariate values and
157 averaged to attain one value per species.

158 **Familiarity** refers to how well-known a species is by the average observer and was quantified by the
159 number of Belgian websites with the Dutch name of the species in the title, retrieved from the Google
160 search engine (Żmihorski et al., 2013). We added two extra search terms that specified the taxonomic
161 group (in Dutch) and excluded the *waarnemingen.be* website to avoid counting individual observations
162 on the used data platform, e.g. "*Bruinrode Heidelibel*" *site:.be libel -waarnemingen.be*. An Incognito
163 window was used to unlink search results from the used google account.

164 **Range size** is the distribution range size of the species during the entire study period 2014-2019 in the
165 study area and was quantified as the total number of grid cells (km²) in which a species has been recorded
166 (McPherson et al., 2004).

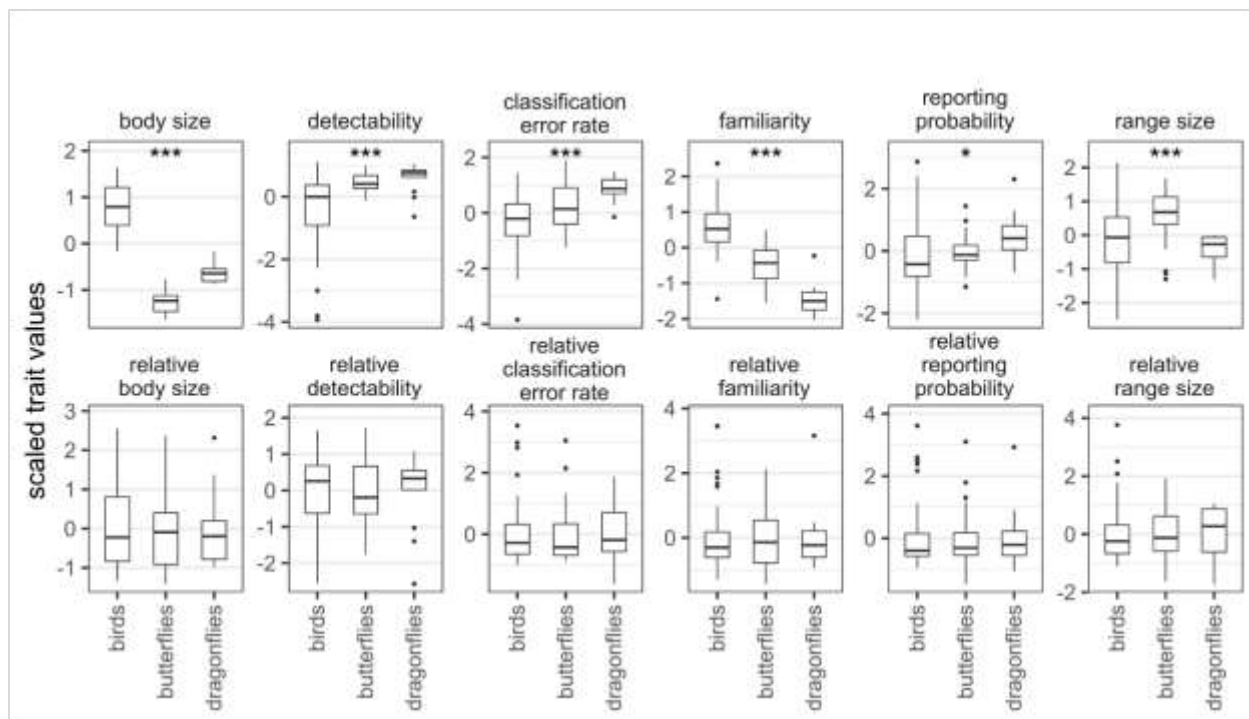
167 **Reporting probability** is the likelihood that a species is reported by an average observer, on the condition
168 that it is present and that the taxonomic group it belongs to is surveyed. To meet these requirements, we
169 looked at the peak of the active season and calculated the relative number of species observations to the
170 number of observations for a taxonomic group. This was averaged across locations and observers. We
171 subsequently divided this number by the average detectability across locations where the species was
172 present to correct for the impact of detectability on reporting rate.

173 **2.3. The impact of data quality filtering**

174 To build recommendations for data quality filtering based on species traits, we first analysed the
175 multivariate relationship between species traits and the filter effects. Consequently, species were
176 grouped in species profiles characterised by the most highly associated traits to assess if such groups
177 presented a similar response to data quality filtering. By filter effect, we mean the impact of data quality
178 filtering by the three filters ACTIVITY (only observations from active observers), DETAIL (only detailed
179 observations) and VALSTAT (only approved observations) on three evaluation metrics: AUC, sensitivity
180 and specificity. All analyses were conducted in R (R Core Team, 2021).

181 **2.3.1. Multi-trait analysis**

182 Relationships between species traits and filter effects were examined using multiple (multi-trait)
183 regressions. The data were modelled in beta-regressions (*betareg* package v3.1-4, Cribari-Neto & Zeileis,
184 2010), because of the bound character of the response variable (Δ AUC, Δ sensitivity and Δ specificity
185 theoretically range from -1 to 1). Filter effect values were rescaled to fall between 0 and 1 with the
186 following transformation: $y = \frac{x - \min(x)}{\max(x) - \min(x)}$. To reduce the impact of outliers, data points with a cook's
187 distance of more than four times the mean cook's distance of all data points were removed (Ferrari et al.,
188 2004). As trait values showed taxonomic differences (Figure 1), continuous values (absolute traits) were
189 rescaled per taxonomic group (relative traits), using the aforementioned transformation (Table 1). The
190 relative values can be informative for patterns across taxonomic groups that would go unnoticed
191 otherwise (e.g. birds are always larger than butterflies, but similar impacts from filtering might be
192 observed for large birds as well as large butterflies).



193
194 *Figure 1: Summary of the species traits per taxonomic group after value transformation and standardisation. Absolute traits (top*
195 *row) were rescaled to relative traits (bottom row) per taxonomic group to assess patterns across taxonomic groups. Stars indicate*
196 *differences in the medians of the trait values between taxonomic groups (***) = $p < 0.001$, * = $p < 0.05$).*

197 First, multi-trait regressions were performed using the log-transformed **absolute** trait values as
198 continuous variables and the taxonomic group as a factor variable. Second, **relative** traits were regressed
199 against the filter effects. Trait values were standardised and multicollinearity was reduced by retaining
200 only those variables with a Variance Inflation Factor (VIF) below 5 (Menard, 2001). We modelled the
201 absolute and relative traits separately because of high pairwise correlations among most of these
202 variables (Figure S1). We also quantified variable importance by leaving out each trait one by one and
203 calculating the decrease in pseudo- R^2 compared to the full model. Finally, we performed a model selection
204 based on three conditions to obtain parsimonious models for each filter effect: (1) the increase in the
205 Akaike's Information Criterion (AIC) had to be smaller than (a conservative) five (Burnham et al., 2011),
206 (2) the model should at least contain the most important variable and (3) the simplest model was selected
207 (i.e. the model with the least parameters).

208 **2.3.2. Species profiles**

209 To test whether groups of species with similar traits can improve recommendations for data quality
210 filtering, we delineated species profiles. Species were clustered into groups with similar traits using the
211 FactoMiner package v1.34 (Le et al., 2008). This package performs a principal component analysis (PCA)
212 on a set of variables (i.e. species traits) followed by an agglomerative hierarchical clustering of individuals
213 (i.e. species) each described by principal components of those variables (Husson et al., 2010). The active
214 variables, included in the PCA and clustering, were those variables that contributed most to the change in
215 model performance across filters, resulting from the multi-trait regression model selection.
216 Supplementary variables were added to characterise the clusters further, without impacting the clustering
217 itself, and comprised: the remaining traits and the impact of filtering on model performance per filter
218 (quantitative), and the taxonomic group (qualitative).

219 We delineated the profiles based on four conditions. First, traits used were those remaining after model
220 selection (see section 2.3.1) in at least one multi-trait regression. Second, the number of clusters was
221 chosen based on the increase of inertia between two consecutive aggregation steps in the hierarchical
222 tree (Husson et al., 2010). Third, profiles were ideally associated with one or more distinctive filter effects:
223 (1) an increase in AUC, (2) a decrease in AUC, (3) an increase in sensitivity and/or decrease in specificity
224 or (4) an increase in specificity and/or decrease in sensitivity. Fourth, profiles should be ecologically
225 meaningful, where we relied on species experts to evaluate the profiles' species composition. To this end,
226 we experimentally changed the number of clusters before selecting the final profiles, by choosing
227 different heights in the hierarchical tree.

228 **2.3.3. Impact of sample size**

229 The role of sample size in the relationship between species traits and filter effects was assessed by adding
230 two sample size situations based on previous recommendations (Van Eupen et al., 2021), where filtering

231 was not advised when sample size was reduced by more than 50% or when the resulting sample size was
232 100 presences. We looked at (1) the **real** situation, i.e. the filter effects when sample size after filtering
233 was not altered (sample size was reduced by an amount that depended on the applied filter), (2) the **r50**
234 situation, i.e. the filter effects when sample size was reduced by 50% or more and (3) the **ss100** situation,
235 i.e. the filter effects when sample size was reduced to 100 presences. Adding these two situations could
236 aid interpretation and simulate situations occurring in datasets of lower quality (i.e. where fewer
237 presences are kept after stringent filtering).

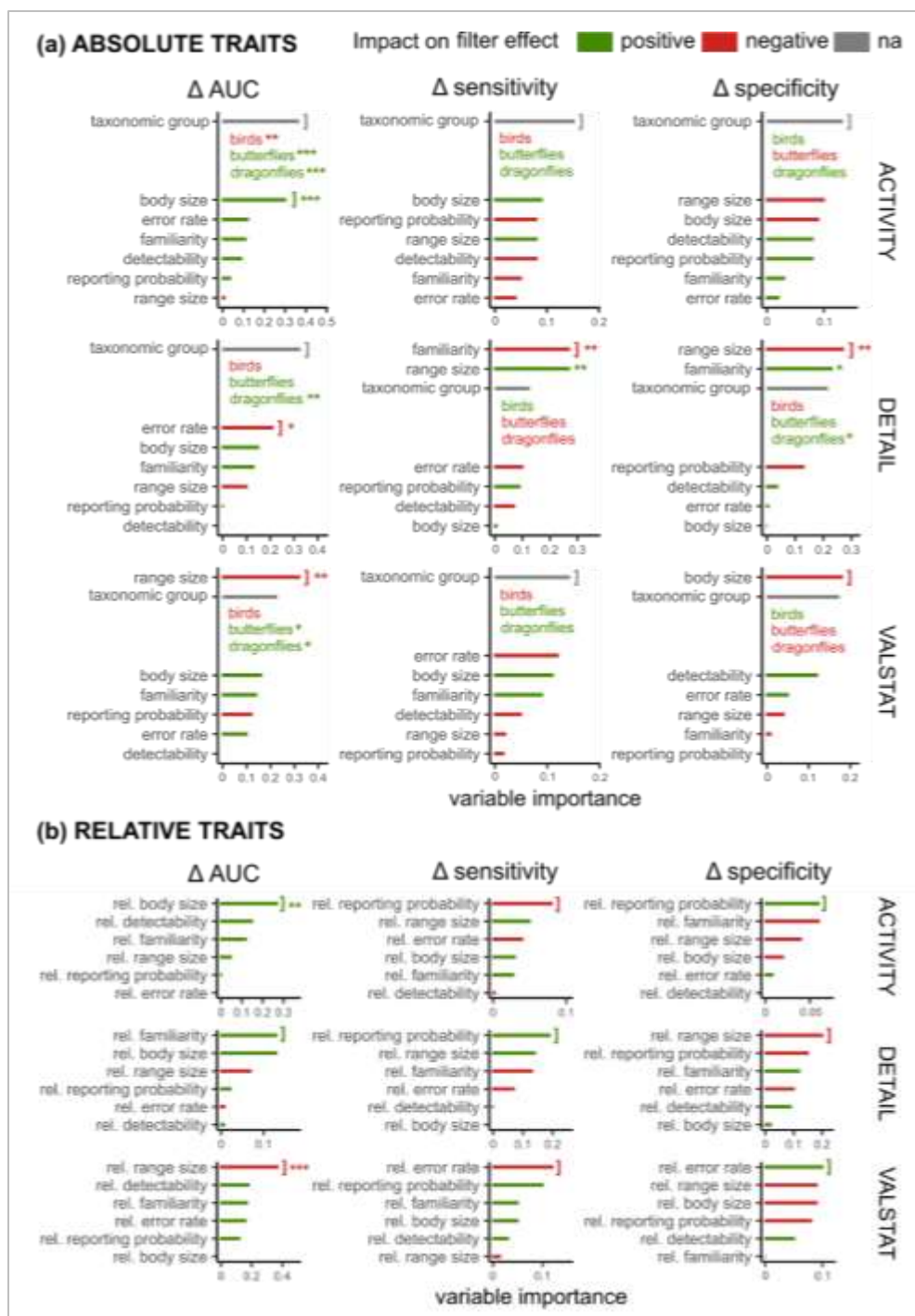
238 **3. Results**

239 **3.1. Multi-trait analysis**

240 Figure 2 shows that the relative importance of the different traits in their association with the filter effects
241 varies among filters and model evaluation metrics. Considering the absolute traits (Figure 2a), the
242 taxonomic group was the most important variable in five out of nine cases. When the goal was to increase
243 AUC, it was best to use data from active observers or approved observations for butterflies and
244 dragonflies, or detailed observations for dragonflies. Specificity could be increased for dragonflies by using
245 detailed observations. No other significant differences ($p < 0.05$) between taxonomic groups were
246 detected in the multiple regressions. AUC of models from large-bodied species could best be increased by
247 using data from active observers, AUC of models from species with a low error rate by using detailed
248 observations and AUC of models from species with a restricted range size by using approved observations.
249 Sensitivity of models from unfamiliar species benefitted from using detailed observations, as did
250 specificity of models from species with restricted range sizes. Specificity of models from small-bodied
251 species benefitted from using approved observations.

252 Considering the relative traits (Figure 2b), using observations from active observers worked best for large-
253 bodied species (to increase AUC), for species with low reporting probability (to increase sensitivity) or for
254 species with high reporting probability (to increase specificity). Using detailed observations most
255 benefitted familiar species (to increase AUC), species with high reporting probability (to increase
256 sensitivity) or species with a restricted range size (to increase specificity). Using approved observations
257 was most valuable for species with a restricted range size (to increase AUC), for species with a low
258 classification error rate (to increase sensitivity) or for species with a high classification error rate (to
259 increase specificity).

260 Multicollinearity among absolute and relative traits was negligible ($VIF < 5$), so all traits were included in
261 the multi-trait regressions. Neither absolute nor relative detectability was retained after model selection
262 as these traits explained less variation compared to others. We did observe that detectability was
263 negatively correlated with familiarity ($r = -0.38$, $p < 0.001$), which can be explained by the taxonomic
264 differences found in both traits (Figure 1 and Figure S1). Detectability was also negatively correlated with
265 reporting probability ($r = -0.62$ and $r = -0.65$ for absolute and relative traits respectively, $p < 0.001$), which
266 can be explained by their inverse dependence (Figure S1 and section 2.2).



267
 268 *Figure 2: Variable importance in the multi-trait regressions for absolute (a) and relative (b) species traits per filter (ACTIVITY,*
 269 *DETAIL and VALSTAT) and change in model evaluation metric (Δ AUC, Δ sensitivity, Δ specificity). Variable importance is expressed*
 270 *as the square root of the change in pseudo- R^2 when leaving out one variable at a time from the full model. Colours indicate a*
 271 *positive (green) or negative (red) impact of the trait on the filter effect, factor variables have grey (n/a) colour. Square brackets*
 272 *indicate the variables kept after model selection (i.e. the simplest model with an increase in the Akaike's Information Criterion*

273 (AIC) of less than 5 compared to the best model where at least the most important variable was included). Asterisks indicate
274 significant model coefficients (*** = $p < 0.001$, ** = $p < 0.01$, * = $p < 0.05$).

275 **3.2. Species profiles**

276 The active variables that we used in the PCA and clustering analysis were the continuous variables kept
277 after model selection in the multi-trait analysis: i.e. body size, relative body size, classification error rate,
278 relative classification error rate, familiarity, relative familiarity, relative reporting probability, range size
279 and relative range size. In the experimental phase, three clusters best captured the variation in species
280 traits, but the impact on model performance still showed large variation within profiles. Ecologically, these
281 profiles also separated species into quite general groups and we cross-checked the clustering of the
282 species into four or more different profiles with species experts. Five clusters appeared the best outcome
283 while keeping cluster size at a reasonable level (minimum cluster size equalled 7 species) (Table 2). The
284 full results of the PCA and clustering analysis are presented in the supplementary information (Table S2
285 and Figure S2).

286 Positive or negative recommendations were only noted when the goal was to increase AUC. When the
287 goal was to increase sensitivity or specificity, recommendations were either cautious or alarming and in
288 most cases, filter recommendations for increasing sensitivity and specificity were opposite to each other.
289 Similar impacts between profiles on one evaluation metric might have a different impact on other metrics
290 (e.g. similar impact on AUC but a different impact on sensitivity and specificity in profiles 1 and 4).

291 *Table 2: Recommendations for data quality filtering for the five species profiles, described by five relative traits (body size,*
292 *classification error rate, familiarity, reporting probability and range size) and four absolute traits (body size, classification error*
293 *rate, familiarity and range size). Recommendations are positive (green – all values in the 90% confidence interval are positive),*
294 *cautious (blue - the average filter effect is positive but the 90% confidence interval also includes negative values), alarming (orange*
295 *- the average filter effect is negative but the 90% confidence interval also includes positive values) or negative (red - all values in*
296 *the 90% confidence interval are negative). The taxonomic distribution of the species is given, as well as the most characterising*

297 *species per profile (in bold are the species closest to the cluster centre and in italic are the species furthest away from the other*
 298 *cluster centres). The asterisks indicate the significance level at which traits, filter effects or taxonomic groups are associated with*
 299 *a profile (** = 0.001, * = 0.01, * = 0.05). For taxonomic groups, (+) and (-) indicate whether the group is significantly more or*
 300 *less represented in a profile.*

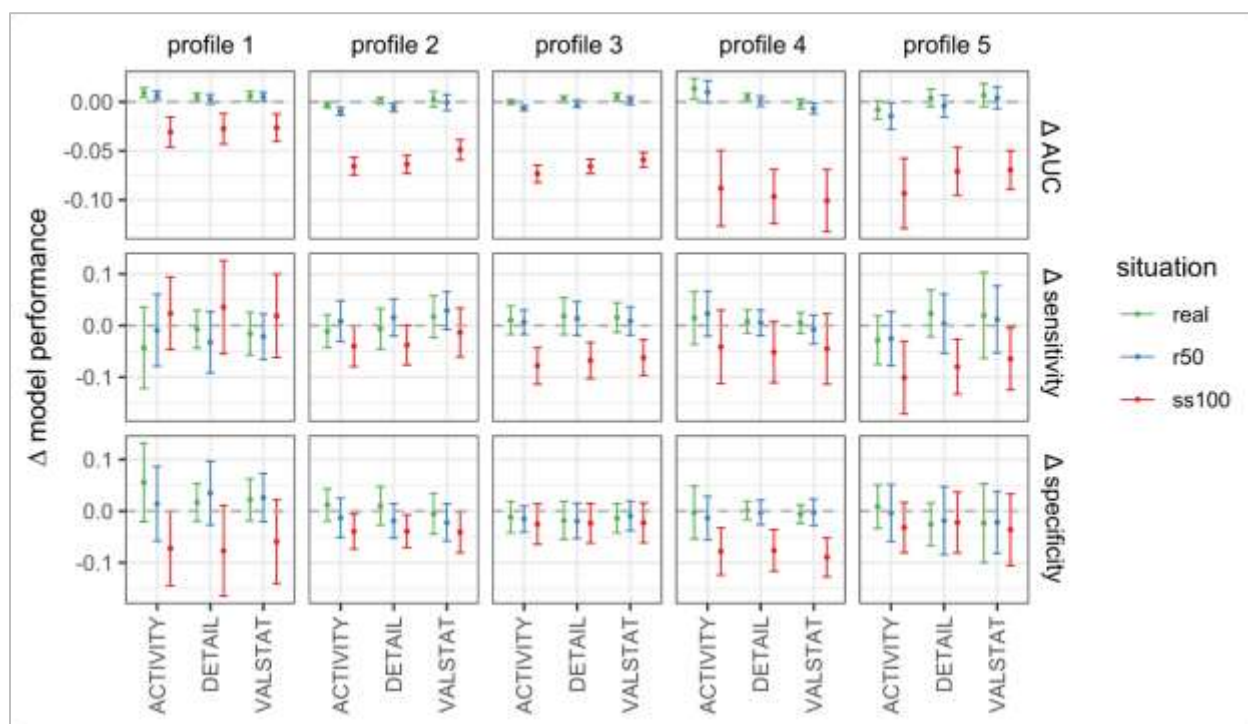
| | PROFILE 1 | PROFILE 2 | PROFILE 3 | PROFILE 4 | PROFILE 5 |
|------------------------------------|--|---|---|--|---|
| <i>Relative Traits</i> | High error rate *** Widespread *** | Small body size *** Restricted range size ** Low error rate * Unfamiliar * | Large body size *** Restricted range size ** Unfamiliar * | Familiar *** Widespread *** Large body size ** Low error rate * | Familiar *** High reporting probability *** Low error rate * |
| <i>AUC recommendations</i> | ACTIVITY ** > VALSTAT > DETAIL | VALSTAT > DETAIL ACTIVITY ** | VALSTAT DETAIL ACTIVITY | ACTIVITY ** > DETAIL VALSTAT | VALSTAT > DETAIL ACTIVITY * |
| <i>sensitivity recommendations</i> | ACTIVITY < VALSTAT < DETAIL | VALSTAT ACTIVITY < DETAIL | DETAIL > VALSTAT > ACTIVITY | ACTIVITY > DETAIL > VALSTAT | DETAIL > VALSTAT ACTIVITY |
| <i>specificity recommendations</i> | ACTIVITY > VALSTAT > DETAIL | ACTIVITY > DETAIL VALSTAT | DETAIL < VALSTAT < ACTIVITY | DETAIL VALSTAT < ACTIVITY | ACTIVITY DETAIL < VALSTAT |
| <i>Taxonomic group</i> | 20 species 4 birds *** (-) 6 butterflies 10 dragonflies *** (+) | 35 species 20 birds 11 butterfly 4 dragonflies | 17 species 17 birds *** (+) | 12 species 4 birds 8 butterflies ** (+) | 7 species 7 birds * (+) |
| <i>Absolute traits</i> | High error rate *** Unfamiliar ** | Restricted range size *** Small body size *** Low error rate * | Large body size *** Low error rate * | Widespread *** Low error rate * | Familiar *** Low error rate * |
| <i>Characterising species</i> | Pieris napi <i>Sympetrum striolatum</i> Sympetrum sanguineum Aeshna cyanea Maniola jurtina <i>Pieris brassicae</i> <i>Pieris rapae</i> <i>Enallagma cyathigerum</i> <i>Larus canus</i> | Oenanthe oenanthe Turdus pilaris Tachybaptus ruficollis Delichon urbicum Rallus aquaticus <i>Platycnemis pennipes</i> <i>Colias crocea</i> <i>Calopteryx splendens</i> <i>Pyrrhosoma nymphula</i> <i>Motacilla alba</i> | Tadorna tadorna Circus aeruginosus Numenius arquata Egretta garzetta Branta leucopsis <i>Cygnus olor</i> <i>Branta canadensis</i> <i>Anser anser</i> <i>Ardea alba</i> <i>Corvus frugilegus</i> | Vanessa cardui Polygonia c.album Gonepteryx rhamni Vanessa atalanta Falco tinnunculus <i>Buteo buteo</i> <i>Aglais io</i> <i>Papilio machaon</i> | Cuculus canorus Alcedo atthis Perdix perdix Carduelis carduelis Athene noctua <i>Ciconia ciconia</i> <i>Alopochen aegyptiaca</i> |

301 Absolute traits appeared highly associated with the most represented taxonomic group for four out of
302 five profiles (Table 2 and Figure 1). Profile 1 contained more dragonflies, indeed species with a higher
303 classification error rate that are less familiar. Profile 4 contained more butterflies, species with a large
304 range size, yet not necessarily a lower error rate. Profile 3 contained birds only, which have larger body
305 sizes and lower error rates. There is a difference with profile 5 though, also containing only birds, where
306 higher familiarity and lower error rates are characterising traits. Profile 2 is not associated with one of the
307 three taxonomic groups, but species in this profile are mostly small, with a restricted range size and a
308 lower error rate, which are also relative traits that characterise this profile.

309 Recommendations based on relative traits were mostly similar to the results in the multi-trait analysis,
310 with a few exceptions (Table 2 and Figure 2). Model AUC for large species increased when using
311 observations from active observers, confirmed by negative and positive recommendations in profiles 2
312 and 4 respectively. In profile 3, however, body size seemed subordinate to the taxonomic group. Higher
313 reporting probability was associated with a higher Δ sensitivity when using detailed observations and with
314 a lower Δ sensitivity when using observations from active observers, confirmed in profile 5. Familiarity
315 had a positive impact on Δ AUC when using detailed observations (DETAIL), confirmed by positive and
316 cautious recommendations in profiles 4 and 5, yet only as the second-best option. Using DETAIL did not
317 necessarily worsen model AUC for unfamiliar species (profile 2 and 3), but not all species in these profiles
318 were unfamiliar (indicated by the weak significance level). For using only approved observations, range
319 size was a good indicator of a change in AUC (profiles 2, 3 and 4), except for the widespread species in
320 profile 1 where range size seemed to be subordinate to error rate. For using only detailed observations,
321 however, range size did not seem to drive filter recommendations when the goal was to increase model
322 specificity, except for the cautious recommendation for species with a restricted range size in profile 2.
323 Finally, the association between error rate and model sensitivity and specificity supported filter
324 recommendations when using only approved observations (profiles 1, 2, 4 and 5).

325 3.3. Impact of sample size

326 Reducing sample size further (beyond the already occurring decrease in sample size after data quality
327 filtering) impacted filter effects both positively and negatively (Figure 3). The impact on AUC was generally
328 negative, but the impact on sensitivity and specificity could in a few cases also be positive. It, therefore,
329 depends on the goal of the study whether reducing sample size further might have a desirable effect. Note
330 that there is usually a trade-off between sensitivity and specificity (when sensitivity increases, specificity
331 usually decreases and the other way around) (Jiménez-Valverde, 2012). Variability in the impact on model
332 performance also increased with decreasing sample size, except for species with a restricted range size
333 (profiles 2 and 3).



334
335 *Figure 3: Recommendations for data quality filtering for the five species profiles in the three sample size situations (real = actual*
336 *reduction in sample size when filtering, r50 = sample size reduced by at least 50%, ss100 = sample size reduced to 100 presences).*
337 *Dots and error bars are the means and 90% confidence intervals for the filter effects.*

338 Reducing sample size further mostly worsened model performance, especially when sample size was
339 reduced to 100 presences where recommendations became alarming or negative in most cases. In our
340 dataset, this meant that sample size was reduced by at least 77% (the lowest unfiltered sample size
341 equalled 432 presences). There were a few exceptions where reducing sample size further did have a
342 positive impact on model performance. For example, recommendations for increasing sensitivity could
343 change from alarming to cautious when reducing sample size over 50% for profile 2 (using observations
344 from active observers or detailed observations) and up till sample size reached 100 presences for profile
345 1 (all filters). Results also showed that model sensitivity was more (and specificity was less) impacted by
346 sample size reduction for profiles with birds only (profiles 3 and 5) compared to profiles with more
347 dragonflies and butterflies (profiles 1 and 4).

348 **3.4. Recommendations for data quality filtering**

349 Recommendations for data quality filtering were built on the various results presented in this article. In
350 general, users of opportunistic records should always pay attention when filtering reduces sample size by
351 more than half of its original size, leading to small sample sizes and we generally advise against filtering
352 when sample size is reduced by more than 75%. We further interpreted the filtering recommendations of
353 the PCA and clustering analysis (Table 2) together with the results of the multi-trait analysis (Figure 2). In
354 the following paragraphs, recommendations are formulated with the aim to increase AUC unless specified
355 otherwise.

356 Results showed that taxonomic group (more than absolute traits) and relative traits formed the best basis
357 for filtering recommendations and when we discuss traits in the following paragraphs, we mean the
358 relative values unless specified otherwise. We recommend using only data from active observers when
359 filtering opportunistic records of large or widespread butterfly and dragonfly species (profiles 1 and 4)
360 and approved observations when filtering bird records unless they are very familiar and widespread

361 (profile 4). In the cases where absolute traits were retained after model selection (Figure 2), it was the
362 relative rather than the absolute trait that was causing the filter effect. For example, dragonflies and
363 butterflies benefitted more from using observations from active observers (ACTIVITY) compared to birds,
364 yet a higher absolute body size also impacted this effect positively. This meant that dragonflies and
365 butterflies with a higher (relative) body size benefitted most from using the ACTIVITY filter. Keeping bird
366 observations from more active observers only was generally not recommended, except for widespread
367 species with a high classification error rate (profile 1).

368 Recommendations based on the taxonomic group seemed to overrule the impact of body size (profile 3)
369 and we advise against using body size as a motivator for filtering bird species data. Recommendations
370 based on the taxonomic group were also superior to the impact of familiarity (profiles 4 and 5), yet we
371 still recommend using more detailed observations (DETAIL) for familiar species, especially when they have
372 high reporting probability and an increase in sensitivity is desired. It must be said that recommendations
373 for using the filter DETAIL showed more inconsistencies compared to the other filters and this filter effect
374 could less clearly be linked to species traits.

375 We recommend using approved observations for species with a restricted range size, especially for large
376 birds. One noted exception was for the widespread species with a high classification error rate (profile 1),
377 where approved observations did impact model AUC positively.

378 When model AUC increased after filtering, sensitivity mostly increased and specificity decreased, with two
379 exceptions noted. First, species in profiles 1 and 2 were generally more difficult to identify, reflected by
380 either a high classification error rate (profile 1) or because they were small-bodied and unfamiliar to an
381 average observer (profile 2). For these profiles, we see that an increase in data quality by using either
382 filter could reduce the impact of false positives on model performance (i.e. increase specificity), except
383 for using approved observations for widespread species in profile 2. A side-effect was that sensitivity had

384 a greater potential to increase when sample size was reduced beyond the real data situation, even at high
385 reductions (Figure 3). A second exception, where specificity increased after filtering, was noted for familiar
386 species when using more detailed observations (profile 4) or data from active observers (profile 5). While
387 the positive impact of using only data from active observers on Δ specificity could be linked to higher
388 reporting probability (profile 5), the positive impact of using only detailed observations for familiar and
389 widespread species contradicted the negative association of Δ specificity with range size (Figure 2).

390

391 **4. Discussion**

392 In this study, we built recommendations for data quality filtering of opportunistic citizen science data
393 when used as input in species distribution models (SDMs), based on a set of a priori defined species traits.
394 Traits associated with a change in model performance after filtering were: body size, classification error
395 rate, familiarity, reporting probability and range size. Based on these traits, it was possible to generate
396 ecologically meaningful species profiles and make filtering recommendations (section 3.4). The analysis
397 of the species profiles mostly agreed with the results of a regression analysis but also gave new insights
398 on the relative importance of the different traits and trait combinations that lead to specific filtering
399 recommendations.

400 One of the main results was that, when choosing a quality filter, the taxonomic group a species belongs
401 to should be considered. This confirms previous findings based on the same dataset (Van Eupen et al.,
402 2021) and makes sense as taxonomic groups by default present differences in most of the considered
403 species traits due to differences in appearance, appeal, distribution etc. In an attempt to simplify the
404 results presented in this study, we have tested different approaches to generate the species profiles:
405 considering relative traits only, clustering of species for each taxonomic group separately and including
406 filter effects as active variables. Unfortunately, none of these approaches lead to profiles that were
407 ecologically more meaningful compared to the profiles suggested here (Table 2 and Table S1; evaluated
408 by species experts). Moreover, they lead to less consistent results (sections 2.3.1 and 2.3.2) or less explicit
409 filtering recommendations (i.e. larger confidence intervals in Figure 3). This confirms the expectation of
410 Van Eupen et al. (2021) who concluded that filtering recommendations can differ between taxonomic
411 groups, but that there might also be common traits among these groups that can refine them. The
412 selected approach indeed revealed that it is possible to formulate recommendations based on taxonomic
413 group and relative traits only (Table 2 and section 3.4). Absolute traits did not directly support
414 recommendations but aided the formation and interpretation of the species profiles as they either

415 characterized the most represented taxonomic group(s) or confirmed a profile's association with relative
416 traits.

417 The taxonomic bias towards bird species in citizen science data could explain some results, as it indicates
418 greater knowledge by the general public of this species group versus other groups such as butterflies and
419 dragonflies (Troudet et al., 2017; <https://waarnemingen.be/stats/>). As increased observer activity can
420 lead to higher experience and expertise (Johnston et al., 2017), this can explain why observer activity
421 mattered more for the less known taxonomic groups in this study (i.e. butterflies and dragonflies). For
422 example, experienced observers were better at detecting individuals of low-density insect populations
423 (Fitzpatrick et al., 2009) and increasing volunteer performance through training could reduce false positive
424 observations for pollinating insects (Ratnieks et al., 2016). These results can also be generalised to other
425 well-known taxonomic groups such as plants. Observer experience, for example, did not increase
426 volunteer performance for identifying an invasive plant species (Crall et al., 2011). Here, observers' self-
427 identified that comfort level was a better predictor of volunteer success.

428 The positive impact of using approved observations for birds, and especially for species with a restricted
429 range size, can be linked to the mechanism of record verification in the database (Swinnen et al., 2018),
430 whereby records that can be verified by photograph or sound play an important role. The verification
431 procedure consists of two main steps: (1) automated record validation by either image recognition or both
432 spatial and temporal proximity of new records to existing approved records and (2) manual expert
433 verification (when there is uncertainty in step 1). A decent photograph or sound record can thus easily
434 lead to multiple approved records and, by consequence, high photo or sound rates have more chance of
435 leading to approved (filtered) datasets of higher quality. Photo rates were, for example, generally higher
436 for bird and butterfly species with restricted range sizes (Table S1), which can explain why they benefitted
437 from using approved records. High photo or sound rates can also reduce the negative impact of locational
438 errors on model performance, especially for small sample sizes (Mitchell et al., 2017). Photographs are

439 often made from a closer distance, especially with the available easy-to-use identification apps (e.g.
440 ObsIdentify), leading to observations with lower locational uncertainty. When they are made from larger
441 distances, mostly for larger species (i.e. birds in this study), smartphone cameras will not suffice and an
442 observer needs a stronger camera lens. We believe that this is a pastime largely practised by more
443 experienced birders that are more likely to correctly register an individual's exact location compared to
444 an inexperienced observer. As for the importance of sound fragments, bird song usually indicates
445 territorial behaviour (Catchpole and Slater, 2008), hence observations made by sound are usually made
446 in birds' respective habitats. Additionally, the prevalence of locational errors in opportunistic bird data
447 will be larger compared to invertebrate species because of their high mobility (Maes et al., 2019), even at
448 a scale of 1 km², which was the resolution used in this study.

449 Large range size is associated with lower model performance because wide-ranged species usually occupy
450 a broad environmental niche and have less distinctive links with their habitat compared to species with a
451 restricted range size that usually have a narrow niche (e.g. Hernandez et al., 2006; Stockwell & Peterson,
452 2002). While increasing model performance for more widespread species through statistical methods or
453 survey design has been observed to be difficult (Brotons et al., 2007; Tassarolo et al., 2014), we observed
454 that using filtered data, especially from more active observers, had a positive impact on model
455 performance for widespread species. We argue, however, that range size in those cases is subordinate to
456 either classification error rate or the taxonomic group. Firstly, improving data quality is always important
457 for any species with a high misidentification risk (Table 2, profile 1). Misidentification errors can distort
458 estimates of species distributions (Costa et al., 2015; Cruickshank et al., 2019; Miller et al., 2011), even
459 though such errors were reduced by spatial aggregation of records (Kramer-Schadt et al., 2013; Van Eupen
460 et al., 2021). Misidentification risk has been found higher for species with similar physical appearance, for
461 example, because they are genetically related (Vantieghem et al., 2017) or have mimicking congeners
462 (Ratnieks et al., 2016). Secondly, widespread species in profile 4 are mostly large butterflies and, as

463 previously discussed, this taxonomic group might benefit more from using data from active observers.
464 Moreover, based on the relative traits only (i.e. not considering the taxonomic group) one would
465 intuitively assume that data quality filtering does not have such a pronounced positive impact in profile 4
466 because these widespread species were also more familiar and had lower error rates.

467 While retaining observations from active observers or approved observations showed clear associations
468 with taxonomic groups or relative species traits, retaining detailed observations showed more
469 inconsistencies, except for the positive impact on model performance for familiar species. Familiarity
470 might reflect the level of detail at which a species' ecology is known, hence data quality can be increased
471 by retaining more detailed observations for species that are familiar to an average citizen scientist.
472 Because retaining only detailed observations on average had the largest impact on sample size (Van Eupen
473 et al., 2021), the impact of sample size may be overruling the effect of the increase in data quality.
474 Reducing the sample size of presences generally impacts presence-only SDMs negatively as model
475 performance decreases, especially at low sample sizes, and performance variability increases (Hernandez
476 et al., 2006; Liu et al., 2019; van Proosdij et al., 2016). An increase in variability was mostly noted for
477 widespread species, as these species are more sensitive to small sample sizes (Liu et al., 2019). While large
478 reductions in sample size require attention, it remains important to realise that filtering simultaneously
479 increases data quality and thus model performance can also increase, especially when less than half of
480 the presences in a dataset are removed (Van Eupen et al., 2021).

481 Detectability did not appear to be an important trait in this study, while it has repeatedly been proven to
482 impact model performance positively (e.g. Pöyry et al., 2008; Seoane et al., 2005), and variation in
483 detectability is directly linked to the problem of imperfect detection in opportunistic presence-only data
484 (Dorazio, 2014). Species traits that are associated with increased detectability are, for example, high
485 abundance (Mccarthy et al., 2013), high singing rates (Sólymos et al., 2018), large body size (Johnston et
486 al., 2014; Pöyry et al., 2008), long lifespan and migratory behaviour (Carrascal et al., 2006). However, we

487 did not find proof that any of these traits were confounded with detectability in our analysis. One trait
488 that could have influenced the outcome for detectability was reporting probability because the way we
489 calculated reporting probability caused a moderate negative correlation between relative reporting
490 probability and relative detectability (Figure S1). However, reporting probability characterised only one
491 profile and thus implications for filtering recommendations would remain marginal.

492 While the highly fragmented (Antrop, 2004) and easily accessible landscape in our study region, Flanders,
493 has many benefits for studying species distributions, it was also one of the limitations. The largest benefit
494 was the consequent high spatial and temporal density of records in the *waarnemingen.be* database
495 (Herremans et al., 2018). On the other hand, because of the high density, the low importance of
496 detectability in our study could be an underestimation when studying regions with less fragmented and
497 larger conservation areas.

498 Another limitation was the insufficient availability of structured data for external model validation in the
499 original dataset (Van Eupen et al., 2021) leading to two restrictive features. First, data consisted of
500 relatively common species (minimum sample size was 432 presences). Rare habitat specialists from
501 habitats with restricted distribution ranges in Flanders (e.g. heathland) were thereby excluded from this
502 analysis. Since these are often targeted species in national and international biodiversity policy (De Ro et
503 al., 2021; Vanden Broeck et al., 2017), it would be useful to adjust the model validation strategy used in
504 Van Eupen et al. (2021) for those species to be able to formulate generic recommendations. Based on the
505 available data, building SDMs with validated data (for species with a restricted range size) or with data
506 from more active observers (for conspicuous invertebrates) could deliver the best results. Second, the
507 data showed sub-optimal representativeness of the taxonomic groups by the studied species. We argue,
508 however, that this imbalance in species representation is often inherent to opportunistic datasets (e.g.
509 over-representation of large birds in Callaghan et al., 2021).

510 Finally, some filter effects might have been impacted by the temporal and spatial aggregation of records
511 over the period 2014-2019 and in grid cells of 1x1 km. While a 1 km² resolution is a standard resolution in
512 Flemish biodiversity studies (e.g. Demolder et al., 2014; Rutten et al., 2019; Vantieghem et al., 2017),
513 performing the analysis at different scales might reveal higher or lower impacts of some traits.

514 **Conclusions**

515 Many have attempted to disentangle the relationships between species ecology and model performance,
516 and this study adds to that knowledge with some basic recommendations for data quality filtering for
517 three commonly studied taxonomic groups. Clustering species in species profiles based on traits that
518 resulted from a multiple regression analysis both highlighted the relative importance of species traits and
519 revealed new insights, and it is important to realise that one single trait does not necessarily predict a
520 species' response to filtering. We found that both the taxonomic group (more than absolute traits) and
521 relative species traits (rescaled values that can be compared among taxonomic groups) defined the impact
522 of data quality filtering on model performance. Our findings largely supported on: (1) the general species
523 knowledge among citizen scientists, with high importance of data quality for widespread and familiar
524 species in general and, more specifically, high importance of observer experience for less known
525 taxonomic groups; and (2) the mechanism of record verification in an opportunistic data platform, with
526 the high importance of submitting observations that can easily be verified, especially for species with
527 restricted range sizes. We encourage the further improvement of general species knowledge and
528 optimisation of record verification protocols in large citizen science projects. While adopting these
529 recommendations, it is always important to keep the goal of the study in mind (i.e. increasing model
530 discrimination capacity, sensitivity and/or specificity) and to keep an eye on the change in sample size
531 caused by stringent filtering.

532 **ACKNOWLEDGEMENTS**

533 We foremost thank the thousands of volunteers for collecting the millions of records that supported this
534 study; *Natuurpunt Studie* for making the data available for this research, in particular Joeri Belis and Karin
535 Gielen for composing the dataset; and Tim Adriaens for his help with the interpretation of the data. We
536 would also like to thank the anonymous reviewers for kindly improving the first version of this manuscript.
537 This work was supported by the Flemish Research Foundation FWO–SB [grant number 1S92118N].

538 **5. References**

- 539 Antrop, M., 2004. Landscape change and the urbanization process in Europe. *Landsc. Urban Plan.* 67, 9–
540 26. [https://doi.org/10.1016/S0169-2046\(03\)00026-4](https://doi.org/10.1016/S0169-2046(03)00026-4)
- 541 Bink, F.A., 1992. *Ecologische atlas van de dagvlinders van Noordwest-Europa*. Schuyt & Co Uitgevers en
542 Importeurs bv, Haarlem.
- 543 Bird, T.J., Bates, A.E., Lefcheck, J.S., Hill, N.A., Thomson, R.J., Edgar, G.J., Stuart-Smith, R.D.,
544 Wotherspoon, S., Krkosek, M., Stuart-Smith, J.F., Pecl, G.T., Barrett, N., Frusher, S., 2014. Statistical
545 solutions for error and bias in global citizen science datasets. *Biol. Conserv.* 173, 144–154.
546 <https://doi.org/10.1016/j.biocon.2013.07.037>
- 547 Brotons, L., Herrando, S., Pla, M., 2007. Updating bird species distribution at large spatial scales:
548 applications of habitat modelling to data from long-term monitoring programs. *Divers. Distrib.* 13,
549 276–288. <https://doi.org/10.1111/j.1472-4642.2007.00339.x>
- 550 Burgess, H.K., Debey, L.B., Froehlich, H.E., Schmidt, N., Theobald, E.J., Ettinger, A.K., Hillerislambers, J.,
551 Tewksbury, J., Parrish, J.K., 2017. The science of citizen science: Exploring barriers to use as a
552 primary research tool. *Biol. Conserv.* 208, 113–120. <https://doi.org/10.1016/j.biocon.2016.05.014>
- 553 Burnham, K.P., Anderson, D.R., Huyvaert, K.P., 2011. AIC model selection and multimodel inference in
554 behavioral ecology: some background, observations, and comparisons. *Behav. Ecol. Sociobiol.* 65,
555 23–35. <https://doi.org/10.1007/s00265-010-1029-6>
- 556 Callaghan, C.T., Poore, A.G.B., Hofmann, M., Roberts, C.J., Pereira, H.M., 2021. Large-bodied birds are
557 over-represented in unstructured citizen science data. *Sci. Rep.* 11, 19073.
558 <https://doi.org/10.1038/S41598-021-98584-7>
- 559 Carrascal, L.M., Javier, S., Palomino, D., Alonso, C.L., Lobo, J.M., 2006. Species-specific features affect the
560 ability of census-derived models to map winter avian distribution. *Ecol. Res.* 21, 681–691.
561 <https://doi.org/10.1007/s11284-006-0173-y>
- 562 Catchpole, C.K., Slater, P.J.B., 2008. *Bird Song*, 2nd ed. ed. Cambridge University Press, Cambridge.
563 <https://doi.org/10.1017/CBO9780511754791>
- 564 Chefaoui, R.M., Lobo, J.M., Hortal, J., 2011. Effects of species' traits and data characteristics on
565 distribution models of threatened invertebrates. *Anim. Biodivers. Conserv.* 34, 229–247.
- 566 Colwell, R.K., Xuan Mao, C., Chang, J., 2004. Interpolating, extrapolating, and comparing incidence-based
567 species accumulation curves. *Ecology* 85, 2717–2727.
- 568 Costa, H., Foody, G., Jiménez, S., Silva, L., 2015. Impacts of Species Misidentification on Species
569 Distribution Modeling with Presence-Only Data. *ISPRS Int. J. Geo-Information* 4, 2496–2518.
570 <https://doi.org/10.3390/ijgi4042496>
- 571 Crall, A.W., Newman, G.J., Stohlgren, T.J., Holfelder, K.A., Graham, J., Waller, D.M., 2011. Assessing
572 citizen science data quality: an invasive species case study. *Conserv. Lett.* 4, 433–442.
573 <https://doi.org/10.1111/J.1755-263X.2011.00196.X>
- 574 Cribari-Neto, F., Zeileis, A., 2010. Beta Regression in R. *J. Stat. Softw.* 34, 1–24.
575 <https://doi.org/10.18637/jss.v034.i02>

- 576 Cruickshank, S.S., Bühler, C., Benedikt, J., Schmidt, R., 2019. Quantifying data quality in a citizen science
577 monitoring program: False negatives, false positives and occupancy trends. *Conserv. Sci. Pract.* 1,
578 e54. <https://doi.org/10.1111/CSP2.54>
- 579 De Ro, A., Vanden broeck, A., Verschaeve, L., Van Dyck, H., Jacobs, I., T’Jollyn, F., Maes, D., 2021.
580 Occasional long distance dispersal does not prevent inbreeding in a threatened butterfly. *BMC*
581 *Ecol. Evol.* 21, 224.
- 582 Demolder, H., Schneiders, A., Spanhove, T., Maes, D., Van Landuyt, W., Adriaens, T., 2014. Hoofdstuk 4
583 Toestand biodiversiteit (INBO.R.2014.6194611), Natuurrapport - Toestand en trends van de
584 ecosystemen en ecosysteemdiensten in Vlaanderen. Instituut voor Natuur- en Bosonderzoek,
585 Brussels.
- 586 Dobson, A.D.M., Milner-Gulland, E.J., Aebischer, N.J., Beale, C.M., Brozovic, R., Coals, P., Critchlow, R.,
587 Dancer, A., Greve, M., Hinsley, A., Ibbett, H., Johnston, A., Kuiper, T., Le Comber, S., Mahood, S.P.,
588 Moore, J.F., Nilsen, E.B., Pocock, M.J.O., Quinn, A., Travers, H., Wilfred, P., Wright, J., Keane, A.,
589 2020. Making Messy Data Work for Conservation. *One Earth* 2, 455–465.
590 <https://doi.org/10.1016/J.ONEAR.2020.04.012>
- 591 Dorazio, R.M., 2014. Accounting for imperfect detection and survey bias in statistical analysis of
592 presence-only data. *Glob. Ecol. Biogeogr.* 23, 1472–1484. <https://doi.org/10.1111/GEB.12216>
- 593 Ferrari, S., Cribari-Neto, F., Ferrari, S.L.P., 2004. Beta Regression for Modelling Rates and Proportions. *J.*
594 *Appl. Stat.* 31, 799–815. <https://doi.org/10.1080/0266476042000214501>
- 595 Fielding, A.H., Bell, J.F., 1997. A review of methods for the assessment of prediction errors in
596 conservation presence/absence models. *Environ. Conserv.* 24, 38–49.
597 <https://doi.org/10.1017/S0376892997000088>
- 598 Fitzpatrick, M.C., Preisser, E.L., Ellison, A.M., 2009. Observer Bias and the Detection of Low-Density
599 Populations. *Ecol. Appl.* 19, 1673–1679.
- 600 Guisan, A., Zimmermann, N.E., 2000. Predictive habitat distribution models in ecology. *Ecol. Modell.* 135,
601 147–186. [https://doi.org/10.1016/S0304-3800\(00\)00354-9](https://doi.org/10.1016/S0304-3800(00)00354-9)
- 602 Hanspach, J., Pompe, S., Klotz, S., 2010. Predictive performance of plant species distribution models
603 depends on species traits. *Perspect. Plant Ecol. Evol. Syst.* 12, 219–225.
604 <https://doi.org/10.1016/j.ppees.2010.04.002>
- 605 Henckel, L., Bradter, U., Jönsson, M., Isaac, N.J.B., Snäll, T., 2020. Assessing the usefulness of citizen
606 science data for habitat suitability modelling: Opportunistic reporting versus sampling based on a
607 systematic protocol. *Divers. Distrib.* 00, 1–15. <https://doi.org/10.1111/ddi.13128>
- 608 Hernandez, P.A., Graham, C.H., Master, L.L., Albert, D.L., 2006. The effect of sample size and species
609 characteristics on performance of different species distribution modeling methods. *Ecography*
610 (Cop.). 29, 773–785.
- 611 Herremans, M., Swinnen, K., Vanreusel, W., Vercaeye, D., Veraghtert, W., Vanormelingen, P., 2018.
612 www.waarnemingen.be. Een veelzijdig portaal voor natuurgegevens. *Natuur.focus* 17, 153–166.
- 613 Husson, F., Josse, J., Pagès, J., 2010. Principal component methods - hierarchical clustering - partitional
614 clustering: why would we need to choose for visualizing data? Applied Mathematics Department.

- 615 Isaac, N.J.B., Jarzyna, M.A., Keil, P., Dambly, L.I., Boersch-Supan, P.H., Browning, E., Freeman, S.N.,
616 Golding, N., Guillera-Aroita, G., Henrys, P.A., Jarvis, S., Lahoz-Monfort, J., Pagel, J., Pescott, O.L.,
617 Schmucki, R., Simmonds, E.G., O'Hara, R.B., 2020. Data Integration for Large-Scale Models of
618 Species Distributions. *Trends Ecol. Evol.* 35, 56–67. <https://doi.org/10.1016/j.tree.2019.08.006>
- 619 Isaac, N.J.B., Pocock, M.J.O., 2015. Bias and information in biological records. *Biol. J. Linn. Soc.* 115, 522–
620 531. <https://doi.org/10.1111/bij.12532>
- 621 Jiménez-Valverde, A., 2012. Insights into the area under the receiver operating characteristic curve
622 (AUC) as a discrimination measure in species distribution modelling. *Glob. Ecol. Biogeogr.* 21, 498–
623 507. <https://doi.org/10.1111/j.1466-8238.2011.00683.x>
- 624 Johnston, A., Fink, D., Hochachka, W.M., Kelling, S., 2017. Estimates of observer expertise improve
625 species distributions from citizen science data. *Methods Ecol. Evol.* 00, 1–10.
626 <https://doi.org/10.1111/2041-210X.12838>
- 627 Johnston, A., Hochachka, W.M., Strimas-Mackey, M.E., Gutierrez, V.R., Robinson, O.J., Miller, E.T., Auer,
628 T., Kelling, S.T., Fink, D., 2021. Analytical guidelines to increase the value of community science
629 data: An example using eBird data to estimate species distributions. *Divers. Distrib.* 27, 1265–1277.
630 <https://doi.org/10.1111/DDI.13271>
- 631 Johnston, A., Newson, S.E., Risely, K., Musgrove, A.J., Massimino, D., Baillie, S.R., Pearce-Higgins, J.W.,
632 2014. Species traits explain variation in detectability of UK birds. *Bird Study* 61, 340–350.
633 <https://doi.org/10.1080/00063657.2014.941787>
- 634 Kamp, J., Oppel, S., Heldbjerg, H., Nyegaard, T., Donald, P.F., 2016. Unstructured citizen science data fail
635 to detect long-term population declines of common birds in Denmark. *Divers. Distrib.* 22, 1024–
636 1035. <https://doi.org/10.1111/ddi.12463>
- 637 Kosmala, M., Wiggins, A., Swanson, A., Simmons, B., 2016. Assessing data quality in citizen science.
638 *Front. Ecol. Environ.* 14, 551–560. <https://doi.org/10.1002/fee.1436>
- 639 Kramer-Schadt, S., Niedballa, J., Pilgrim, J.D., Schröder, B., Lindenborn, J., Reinfelder, V., Stillfried, M.,
640 Heckmann, I., Scharf, A.K., Augeri, D.M., Cheyne, S.M., Hearn, A.J., Ross, J., Macdonald, D.W.,
641 Mathai, J., Eaton, J., Marshall, A.J., Semiadi, G., Rustam, R., Bernard, H., Alfred, R., Samejima, H.,
642 Duckworth, J.W., Breitenmoser-Wuersten, C., Belant, J.L., Hofer, H., Wilting, A., 2013. The
643 importance of correcting for sampling bias in MaxEnt species distribution models. *Divers. Distrib.*
644 19, 1366–1379. <https://doi.org/10.1111/DDI.12096>
- 645 Le, S., Josse, J., Husson, F., 2008. FactoMineR: An R Package for Multivariate Analysis. *J. Stat. Softw.* 25,
646 1–18.
- 647 Lindenmayer, D., Woinarski, J., Legge, S., Southwell, D., Lavery, T., Robinson, N., Scheele, B., Wintle, B.,
648 2020. A checklist of attributes for effective monitoring of threatened species and threatened
649 ecosystems. *J. Environ. Manage.* 262, 110312. <https://doi.org/10.1016/j.jenvman.2020.110312>
- 650 Liu, C., Newell, G., White, M., 2019. The effect of sample size on the accuracy of species distribution
651 models: considering both presences and pseudo-absences or background sites. *Ecography (Cop.)*.
652 42, 535–548. <https://doi.org/10.1111/ecog.03188>
- 653 Lobo, J.M., Jiménez-valverde, A., Real, R., 2008. AUC: A misleading measure of the performance of
654 predictive distribution models. *Glob. Ecol. Biogeogr.* 17, 145–151. <https://doi.org/10.1111/j.1466-8238.2007.00358.x>

- 656 MacKenzie, D.I., Nichols, J.D., Royle, J.A., Pollock, K.H., Bailey, L., Hines, J.E., 2017. *Occupancy Estimation and Modeling: Inferring Patterns and Dynamics of Species Occurrence*. Elsevier.
- 657
- 658 Maes, D., Bauwens, D., De Bruyn, L., Anselin, A., Vermeersch, G., Van Landuyt, W., De Knijf, G., Gilbert, M., 2005. Species richness coincidence: Conservation strategies based on predictive modelling. *Biodivers. Conserv.* 14, 1345–1364. <https://doi.org/10.1007/S10531-004-9662-X>
- 659
- 660
- 661 Maes, D., Brosens, D., Desmet, P., Piesschaert, F., Van Hoey, S., Adriaens, T., Dekoninck, W., Devos, K., Lock, K., Onkelinx, T., Packet, J., 2019. A database of threat statuses and life-history traits of Red List species in Flanders (northern Belgium). *Biodivers. Data J.* 7, e34089. <https://doi.org/10.3897/BDJ.7.e34089>
- 662
- 663
- 664
- 665 Maes, D., Isaac, N.J.B., Harrower, C.A., Collen, B., van Strien, A.J., Roy, D.B., 2015. The use of opportunistic data for IUCN Red List assessments. *Biol. J. Linn. Soc.* 115, 690–706. <https://doi.org/10.1111/bij.12530>
- 666
- 667
- 668 Matutini, F., Baudry, J., Pain, G., Sineau, M., Pithon, J., 2021. How citizen science could improve species distribution models and their independent assessment. *Ecol. Evol.* 11, 3028–3039. <https://doi.org/10.1002/ece3.7210>
- 669
- 670
- 671 McCarthy, M.A., Moore, J.L., Morris, W.K., Parris, K.M., Garrard, G.E., Vesk, P.A., Rumpff, L., Giljohann, K.M., Camac, J.S., Bau, S.S., Friend, T., Harrison, B., Yue, B., 2013. The influence of abundance on detectability. *Oikos* 122, 717–726. <https://doi.org/10.1111/j.1600-0706.2012.20781.x>
- 672
- 673
- 674 McPherson, J.M., Jetz, W., Rogers, D.J., 2004. The effects of species' range sizes on the accuracy of distribution models: Ecological phenomenon or statistical artefact? *J. Appl. Ecol.* 41, 811–823. <https://doi.org/10.1111/J.0021-8901.2004.00943.X>
- 675
- 676
- 677 Menard, S., 2001. *Applied Logistic Regression Analysis*. 2nd edition. SAGE Publications, Inc.
- 678
- 679
- 680
- 681 Mitchell, P.J., Monk, J., Laurenson, L., 2017. Sensitivity of fine-scale species distribution models to locational uncertainty in occurrence data across multiple sample sizes. *Methods Ecol. Evol.* 8, 12–21. <https://doi.org/10.1111/2041-210X.12645>
- 682
- 683
- 684 Morton, E.S., 1975. Ecological Sources of Selection on Avian Sounds. *Am. Nat.* 109, 17–34.
- 685
- 686
- 687
- 688
- 689
- 690
- 691
- 692
- 693
- 694
- 695

- 696 8238.2007.00373.x
- 697 R Core Team, 2021. R: A language and environment for statistical computing.
- 698 Ratnieks, F.L.W., Schrell, F., Sheppard, R.C., Brown, E., Bristow, O.E., Garbuzov, M., 2016. Data reliability
699 in citizen science: learning curve and the effects of training method, volunteer background and
700 experience on identification accuracy of insects visiting ivy flowers. *Methods Ecol. Evol.* 7, 1226–
701 1235. <https://doi.org/10.1111/2041-210X.12581>
- 702 Rutten, A., Casaer, J., Swinnen, K.R.R., Herremans, M., Leirs, H., 2019. Future distribution of wild boar in
703 a highly anthropogenic landscape: Models combining hunting bag and citizen science data. *Ecol.
704 Modell.* 411, 108804. <https://doi.org/10.1016/j.ecolmodel.2019.108804>
- 705 Seoane, J., Carrascal, L.M., Alonso, L., Palomino, D., 2005. Species-specific traits associated to prediction
706 errors in bird habitat suitability modelling. *Ecol. Modell.* 185, 299–308.
707 <https://doi.org/10.1016/j.ecolmodel.2004.12.012>
- 708 Serra-Díaz, J.M., Enquist, B.J., Maitner, B., Merow, C., Svenning, J.C., 2017. Big data of tree species
709 distributions: how big and how good? *For. Ecosyst.* 4. <https://doi.org/10.1186/s40663-017-0120-0>
- 710 Sólymos, P., Matsuoka, S.M., Stralberg, D., Barker, N.K.S., Bayne, E.M., 2018. Phylogeny and species
711 traits predict bird detectability. *Ecography (Cop.)*. 41, 1595–1603.
712 <https://doi.org/10.1111/ecog.03415>
- 713 Steen, V.A., Elphick, C.S., Tingley, M.W., 2019. An evaluation of stringent filtering to improve species
714 distribution models from citizen science data. *Biodivers. Res.* 25, 1857–1869.
715 <https://doi.org/10.1111/ddi.12985>
- 716 Stockwell, D.R.B., Peterson, A.T., 2002. Effects of sample size on accuracy of species distribution models.
717 *Ecol. Modell.* 148, 1–13.
- 718 Storchová, L., Hořák, D., 2018. Life-history characteristics of European birds. *Glob. Ecol. Biogeogr.* 27,
719 400–406. <https://doi.org/10.1111/geb.12709>
- 720 Sullivan, B.L., Wood, C.L., Iliff, M.J., Bonney, R.E., Fink, D., Kelling, S., 2009. eBird: A citizen-based bird
721 observation network in the biological sciences. *Biol. Conserv.* 142, 2282–2292.
722 <https://doi.org/10.1016/j.biocon.2009.05.006>
- 723 Swinnen, K.R.R., Jacobs, A., Claus, K., Ruyts, S., Vercayie, D., Lambrechts, J., Herremans, M., n.d. ‘Animals
724 under wheels’: wildlife roadkill data collection by citizen scientists as a part of their nature
725 recording activities. *Nat. Conserv.*
- 726 Swinnen, K.R.R., Vercayie, D., Vanreusel, W., Barendse, R., Boers, K., Bogaert, J., Dekeukeleire, D.,
727 Driessens, G., Dupriez, P., Jooris, R., Steeman, R., van Asten, K., van den Neucker, T., van
728 Dorsselaer, P., van Vooren, P., Wysmantel, N., Gielen, K., Desmet, P., Herremans, M., 2018.
729 Waarnemingen.be-Non-native plant and animal occurrences in Flanders and the Brussels Capital
730 Region, Belgium. *BiolInvasions Rec.* 7, 335–342. <https://doi.org/10.3391/bir.2018.7.3.17>
- 731 Tessarolo, G., Rangel, T.F., Araújo, M.B., Hortal, J., 2014. Uncertainty associated with survey design in
732 Species Distribution Models. *Divers. Distrib.* 20, 1258–1269. <https://doi.org/10.1111/ddi.12236>
- 733 Thomaes, A., Kervyn, T., Maes, D., 2008. Applying species distribution modelling for the conservation of
734 the threatened saproxylic Stag Beetle (*Lucanus cervus*). *Biol. Conserv.* 141, 1400–1410.

- 735 <https://doi.org/10.1016/j.biocon.2008.03.018>
- 736 Troudet, J., Grandcolas, P., Blin, A., Vignes-Lebbe, R., Legendre, F., 2017. Taxonomic bias in biodiversity
737 data and societal preferences. *Sci. Rep.* 7, 9132. <https://doi.org/10.1038/s41598-017-09084-6>
- 738 Urban, M.C., Bocedi, G., Hendry, A.P., Mihoub, J.B., Pe'er, G., Singer, A., Bridle, J.R., Crozier, L.G., De
739 Meester, L., Godsoe, W., Gonzalez, A., Hellmann, J.J., Holt, R.D., Huth, A., Johst, K., Krug, C.B.,
740 Leadley, P.W., Palmer, S.C.F., Pantel, J.H., Schmitz, A., Zollner, P.A., Travis, J.M.J., 2016. Improving
741 the forecast for biodiversity under climate change. *Science* (80-.). 353, aad8466.
742 <https://doi.org/10.1126/science.aad8466>
- 743 Van Eupen, C., Maes, D., Herremans, M., Swinnen, K.R.R., Somers, B., Luca, S., 2021. The impact of data
744 quality filtering of opportunistic citizen science data on species distribution model performance.
745 *Ecol. Modell.* 444, 109453. <https://doi.org/10.1016/j.ecolmodel.2021.109453>
- 746 van Proosdij, A.S.J., Sosef, M.S.M., Wieringa, J.J., Raes, N., 2016. Minimum required number of specimen
747 records to develop accurate species distribution models. *Ecography* (Cop.). 39, 542–552.
748 <https://doi.org/10.1111/ECOG.01509>
- 749 Van Strien, A.J., Van Swaay, C.A.M., Termaat, T., 2013. Opportunistic citizen science data of animal
750 species produce reliable estimates of distribution trends if analysed with occupancy models. *J.*
751 *Appl. Ecol.* 50, 1450–1458. <https://doi.org/10.1111/1365-2664.12158>
- 752 Vanden Broeck, A., Maes, D., Kelager, A., Wynhoff, I., WallisDeVries, M.F., Nash, D.R., Oostermeijer,
753 J.G.B., Van Dyck, H., Mergeay, J., 2017. Gene flow and effective population sizes of the butterfly
754 *Maculinea alcon* in a highly fragmented, anthropogenic landscape. *Biol. Conserv.* 209, 89–97.
755 <https://doi.org/10.1016/j.biocon.2017.02.001>
- 756 Vantieghem, P., Maes, D., Kaiser, A., Merckx, T., 2017. Quality of citizen science data and its
757 consequences for the conservation of skipper butterflies (Hesperiidae) in Flanders (northern
758 Belgium). *J. Insect Conserv.* 21, 451–463. <https://doi.org/10.1007/s10841-016-9924-4>
- 759 Vermeersch, G., Devos, K., Driessens, G., Evereaert, J., Feys, S., Herremans, M., Onkelinx, T., Stienen,
760 E.W.M., T’Jollyn, F., Anselin, A., 2020. Broedvogels in Vlaanderen 2013-2018. Medelingen van het
761 Instituut voor Natuur- en Bosonderzoek 2020 (1), Brussel.
762 <https://doi.org/10.21436/inbor.18794135>
- 763 Żmihorski, M., Dziarska-Patac, J., Sparks, T.H., Tryjanowski, P., 2013. Ecological correlates of the
764 popularity of birds and butterflies in Internet information resources. *Oikos* 122, 183–190.
765 <https://doi.org/10.1111/J.1600-0706.2012.20486.X>
- 766 Zurell, D., Franklin, J., König, C., Bouchet, P.J., Dormann, C.F., Elith, J., Fandos, G., Feng, X., Guillera-
767 Arroita, G., Guisan, A., Lahoz-Monfort, J.J., Leitão, P.J., Park, D.S., Peterson, A.T., Rapacciuolo, G.,
768 Schmatz, D.R., Schröder, B., Serra-Diaz, J.M., Thuiller, W., Yates, K.L., Zimmermann, N.E., Merow,
769 C., 2020. A standard protocol for reporting species distribution models. *Ecography* (Cop.). 43,
770 1261–1277. <https://doi.org/10.1111/ECOG.04960>
- 771 [dataset] Van Eupen, C., Maes, D., Herremans, M., Swinnen, K.R.R., Somers, B., Luca, S., 2021b. The
772 impact of data quality filtering of opportunistic citizen science data on species distribution model
773 performance: dataset used for Maxent modelling., Dryad, Dataset,
774 <https://doi.org/10.5061/dryad.jwstqjq83>
- 775